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- So far, we have dealt with the Lee & Carter (1992) model and some of its variants. We now look into the AG2022 Model and COVID-19.
- The AG2022 Model consists of three layers (a so-called three-layer Lee-Li model):
 - A layer for the European population.
 - A correction layer for the Dutch population.
 - An extra layer for COVID-19.
- The first two layers are as in AG2020.

	Fixed Age effect Europe	Dynamic Age effect Europe	Fixed Age effect Deviation NL	Dynamic Age effect Deviation NL	Dynamic Age effect COVID in NL
	↓	↓	↓	↓	↓
$\ln(\mu_{x,t}^g) = A_x^g + B_x^g K_t^g + \alpha_x^g + \beta_x^g \kappa_t^g + \tilde{\mathfrak{B}}_x^g \mathfrak{x}_t^g$					
		↑		↑	↑
		Time effect Europe		Time effect Deviation NL	Time effect COVID in NL

- The AG2022 directly builds upon AG2020.
- Surrounding countries selected based upon comparable welfare level.
- AG2020 re-estimated with data EU 2019 added (but data 2020 & 2021 not included).
 - EU data: $D_{x,t}^{(g),EU}, E_{x,t}^{(g),EU}, t = 1970, \dots, 2019, x = 0, \dots, 90$
 - Dutch data: $D_{x,t}^{(g),NL}, E_{x,t}^{(g),NL}, t = 1970, \dots, 2019, x = 0, \dots, 90$
- AG2020 is basis for long-term projections.
- The third layer is due to COVID-19 and is new. It is calibrated using data from 2020 and 2021.

- Number of deaths ($D_{x,t}^{(g)}$) modeled assuming a Poisson-distribution with expectation equal to exposure ($E_{x,t}^{(g)}$) times force of mortality ($\mu_{x,t}^{(g),pre-covid}$):

$$D_{x,t}^{(g)} \mid E_{x,t}^{(g)} \sim \mathcal{P}\left(E_{x,t}^{(g)} \mu_{x,t}^{(g),pre-covid}\right)$$

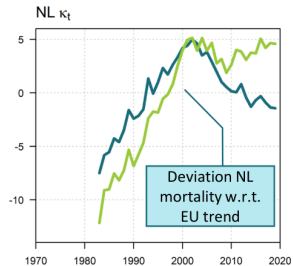
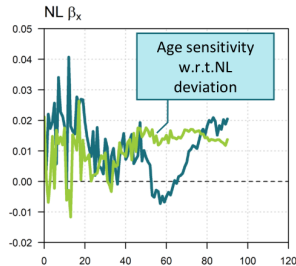
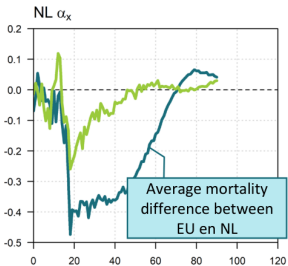
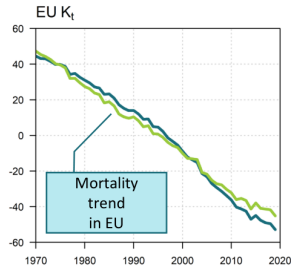
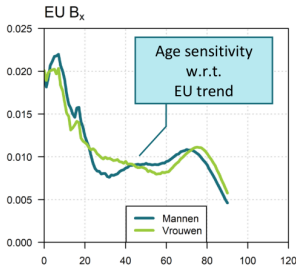
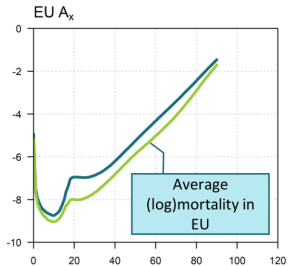
- The long-term mortality trend is estimated based on the European reference group:

$$\ln \mu_{x,t}^{(g),pre-covid,EU} = A_x^{(g)} + B_x^{(g)} K_t^{(g)}$$

- For the Netherlands mortality is estimated conditional on the force of mortality of the European reference group:

$$\ln \mu_{x,t}^{(g),pre-covid} = A_x^{(g)} + B_x^{(g)} K_t^{(g)} + \alpha_x^{(g)} + \beta_x^{(g)} \kappa_t^{(g)}$$

- All parameters are estimated by maximizing the corresponding likelihood functions.



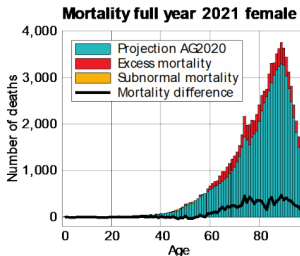
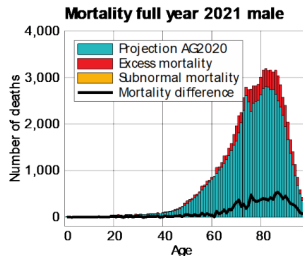
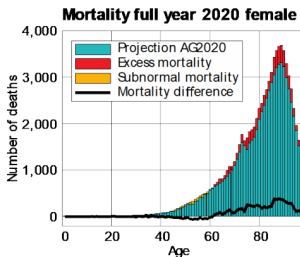
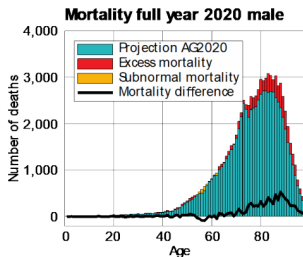
- The European period effects, $K_t^{(g)}$, are assumed to follow a random walk with drift (as in Lee-Carter):

$$\Delta K_t^{(g)} = \theta^{(g)} + \varepsilon_t^{(g)}$$

- The Dutch period effects, $\kappa_t^{(g)}$, are assumed to follow a first-order autoregressive process with constant:

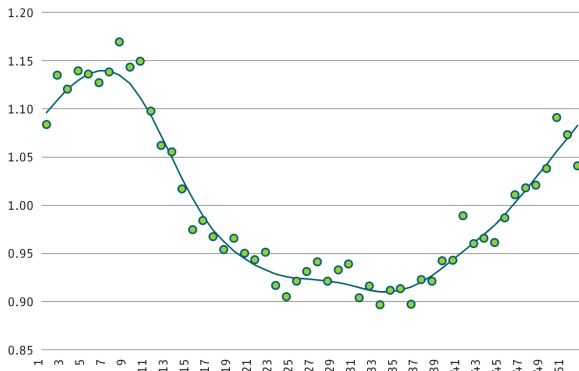
$$\Delta \kappa_t^{(g)} = a^{(g)} \kappa_{t-1}^{(g)} - c^{(g)} + \delta_t^{(g)}$$

- The error terms are assumed to follow a multivariate normal distribution with mean vector 0 and covariance matrix C .
- This combination of assumptions results in coherent projections:
 - In the long run the difference between the Dutch and the European death probabilities converges to zero, but:
 - In the short run the Dutch and European death probabilities might deviate.



- Clear impacts of COVID-19 in both 2020 and 2021.
- Age effect in AG2022 differs from age effect in AG2020.

- Data from 2020 and 2021 cannot be used for normal update.
- Change to weekly data:
 - Customized data needed, plus interpolation and extrapolation.
 - Seasonal effect.
 - New age- and time effect.

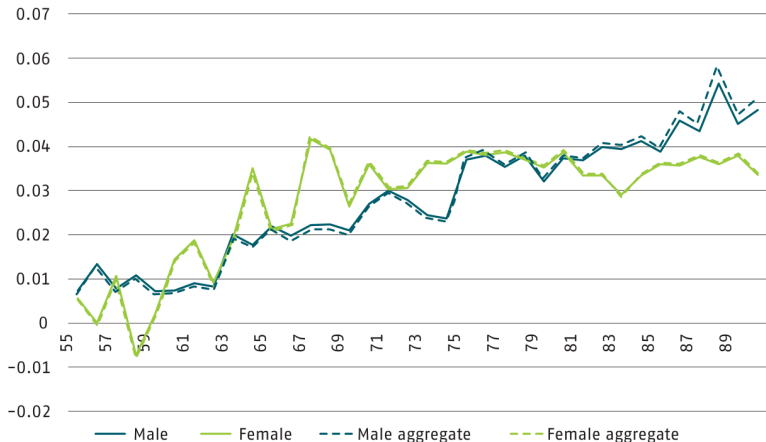


- To describe the deviation in mortality in 2020 and 2021 relative to what is expected on the basis of data before 2020, we introduce a new age effect $\mathfrak{B}_x^{(g)}$ and new (week-based) timeseries $\mathfrak{K}_{w,2020}^{(g)}$ for 2020 and $\mathfrak{K}_{w,2021}^{(g)}$ for 2021.
- The seasonal effect determined earlier is also added.
- Three factors:
 - the pre-pandemic estimate for a given year (an AG2020 model update with datapoints until 2019),
 - a seasonal effect for a given week,
 - a new factor representing the impact of the altered circumstances since the pandemic:

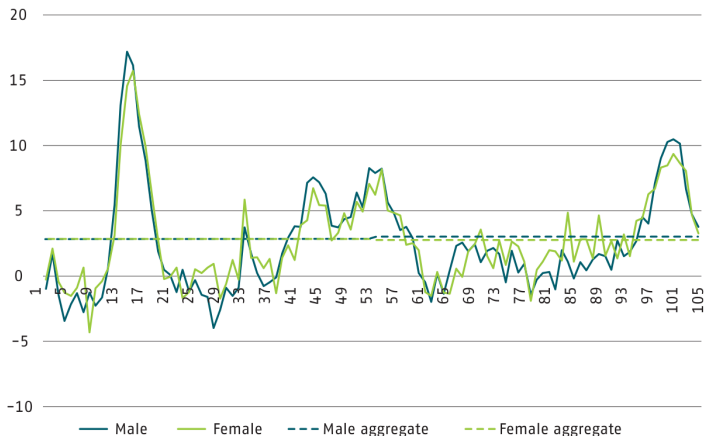
$$\ln \mu_{x,t}^{(g)} = \ln \mu_{x,t}^{(g),pre-covid} + \ln \phi_{w,t} + \mathfrak{B}_x^{(g)} \mathfrak{K}_{w,t}^{(g)}$$

- The time effect of COVID-19 $\mathfrak{K}_{w,t}^{(g)}$ varies significantly across weeks.

Third Layer: Age Effect



- Using weekly data allows the estimation of age effects based on more than 2 (annual) data points.
- No effect until age 55, constant from age 90.



- Time series for difference in mortality (whether or not due to COVID) relative to AG2020 model (with update 2019 data).
- Aggregating week effects $\mathcal{R}_{w,t}^{(g)}$ gives estimate of the impact $\mathcal{X}_t^{(g)}$ for the whole year (dashed lines).

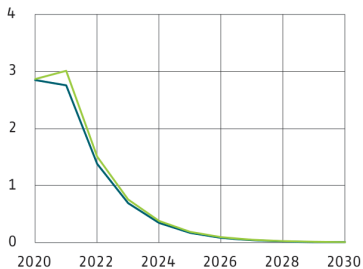
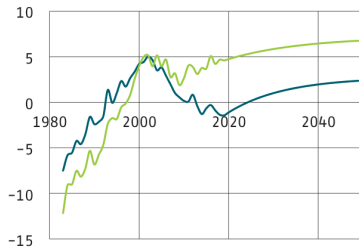
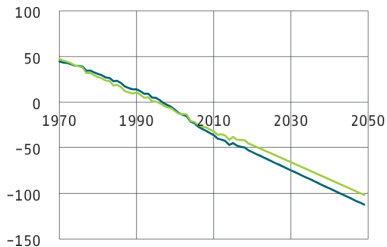
- Estimating the impact for a whole year is the starting point for projecting future years.
- For the course of the projection further assumptions are needed.
- Choice CSO: disappearing (exponentially), i.e.,

$$\mathfrak{X}_t^{(g)} = \mathfrak{X}_{2021}^{(g)} \eta^{t-2021}, \quad t \geq 2021$$

with $\eta = 1/2$, which implies that the half-life of the impact equals one year.

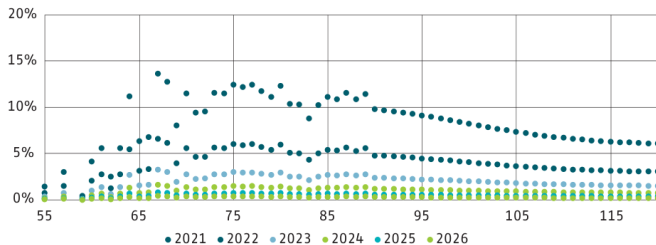
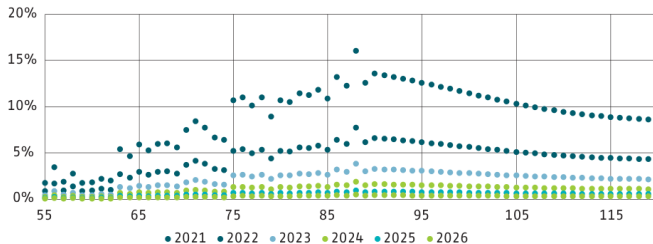
- This assumption determines the 'best estimates' for all future values of the time series in the model.

Third Layer: Projection Model



— Males
— Females

Third Layer: Projection Model

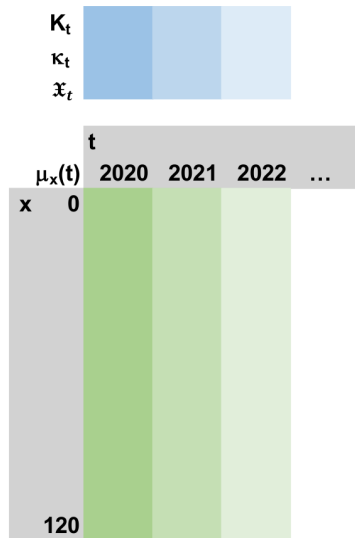
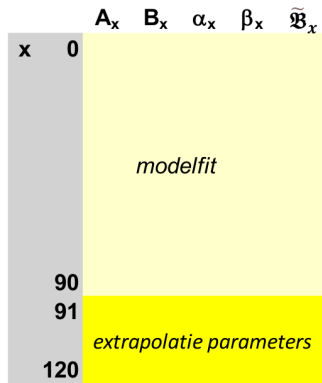


- *Problem:* The AG2022 model uses mortality data for $x = 0, \dots, 90$. If $x \geq 91$, the data becomes rather noisy, because the population older than 90 is small.
- To deal with this issue, the AG2020 model used a popular method to extrapolate mortality rates for older ages – the *Kannisto closing method*.
- For the ages $x \in \{91, \dots, 120\}$, $\hat{\mu}_{x, T+t}^{(g)}$ is determined as

$$\hat{\mu}_{x, T+t}^{(g)} = L\left(\sum_{k=80}^{90} w_k(x) L^{-1}(\hat{\mu}_{k, T+t}^{(g)})\right),$$

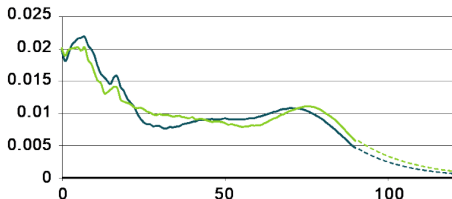
where $L(z) = \frac{1}{1+e^{-z}}$ and $L^{-1}(z) = \ln\left(\frac{z}{1-z}\right)$ and some weight function w_k (see Assingment).

Closure of the Life Table



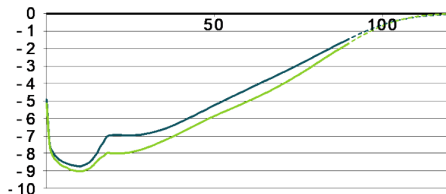
- The Kannisto method was standard in the AG-models and has been standard in many other forecasting models to deal with sparse data points in older cohorts.
- This method has some drawbacks:
 - Not all ages show improved mortality over time. For higher ages mortality rises monotonously to a positive and limit value that is known with certainty, $\lim_{t \rightarrow \infty} \hat{q}_{x, T+t}^{(m)} = 0.6321$, $x \geq 101$. As a result, life expectancy also converges to a limit value known with certainty $\lim_{t \rightarrow \infty} e_{0,t}^{(m)} = 102.08$.
 - Because the limit values are known with certainty, uncertainty decreases (smaller confidence intervals) over time, while we expect increasing uncertainty (wider confidence intervals).
- The new method in AG2022 addresses this issue and closes the life table without knowing limit values with certainty.

$$B_x^g$$



- Extrapolate $\ln(B_x^g)$ linearly
- Interpretation: all ages benefit from decreasing trend in K_t

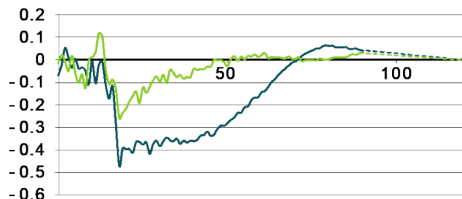
$$A_x^g$$



- Determine A_x^g such that in the last sample period (2019) the same EU death probabilities result as in case of Kannisto

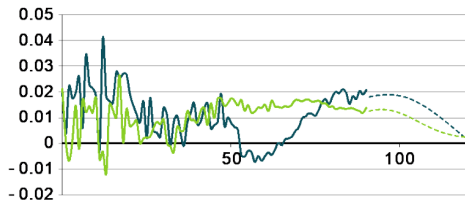
— male - - - - extrapolation
— female - - - - extrapolation

$$\alpha_x^g$$



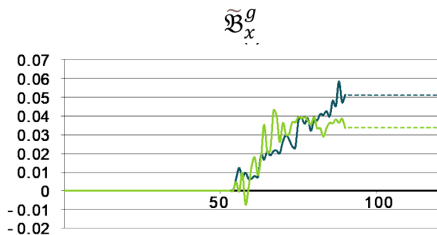
- Extrapolate α_x^g linearly to $\alpha_{120}^g = 0$
- Interpretation:
for age 120 no difference between NL and EU.

$$\beta_x^g$$



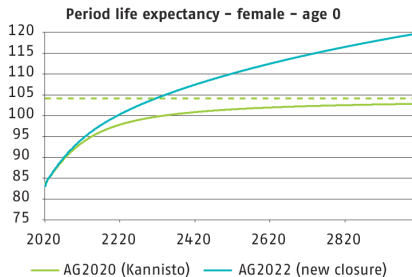
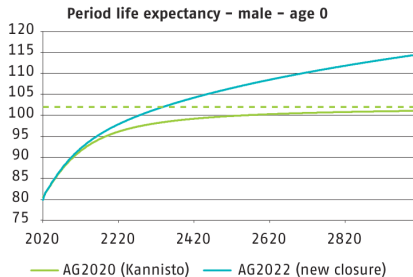
- Determine β_x^g such that in the last sample period (2019) the same NL death probabilities result as in case of Kannisto.





- Parameter COVID-term constant from age 90 on

— male - - - - extrapolation
— female - - - - extrapolation



- The new closing method does not have the undesirable characteristics, which we observe in the application of Kannisto per projection year.
- There is no turning point at a certain age: also, at higher ages the death probabilities keep decreasing.
- The life expectancy does not have a limit whose value is known in advance.
- The confidence intervals do not become smaller over time.

